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## ADVERSARIAL SMOOTHED ANALYSIS

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**Abstract.** The purpose of this note is to extend the results on uniform smoothed analysis of condition numbers from [1] to the case where the perturbation follows a radially symmetric probability distribution. In particular, we will show that the bounds derived in [1] still hold in the case of distributions whose density has a singularity at the center of the perturbation, which we call *adversarial*.

**AMS subject classifications.** Primary 65Y20; secondary 65G99.

**Key words.** Condition numbers, random matrices, average case analysis, smoothed analysis.

**1. Introduction.** Condition numbers play a central role in numerical analysis. They occur in error analysis for finite-precision algorithms (this being historically the reason for their introduction in the late 1940's by von Neumann and Goldstine [10] and Turing [9]) as well as a parameter in expressions bounding the number of iterations in a variety of algorithms (a paradigmatic example being the conjugate gradient method [8, Theorem 38.5]). In practice, however, a difficulty appears: it would seem that to know the condition number of a given data one needs to solve the problem at hand on this data. An inconvenient circularity. A way out of it, proposed by Steve Smale (see [5] for a review), is to assume a probability measure on the space of data and to study the condition number  $\mathcal{C}(a)$  at data  $a$  as a random variable. In other words, to study the condition number of random data.

In doing so Demmel [2] noticed that most condition numbers could be written as (or at least reasonably sharply bounded by) the relativized inverse of the distance from the data  $a \in \mathbb{R}^{n+1}$  to a set of ill-posed instances  $\Sigma \subset \mathbb{R}^{n+1}$ . That is, one could write

$$\mathcal{C}(a) = \frac{\|a\|}{\text{dist}(a, \Sigma)}. \quad (1.1)$$

The simplest example of this phenomenon is given by the condition number for matrix inversion and linear equation solving. For a non-singular square matrix  $A$  it takes the form  $\kappa(A) := \|A\| \|A^{-1}\|$ , where  $\|\cdot\|$  denotes the operator norm. The Condition Number Theorem by Eckart and Young states that  $\|A^{-1}\| = d(A, \Sigma)^{-1}$ , where  $\Sigma$  is the set of singular matrices.

In most applications,  $\Sigma$  is a pointed cone. Therefore, one could normalize so that  $a$  belongs to the  $n$ -dimensional unit sphere  $S^n$ . Note that the usual assumption that  $a$  has a Gaussian distribution in  $\mathbb{R}^{n+1}$  yields a uniform distribution in  $S^n$  after this normalization. It is for condition numbers as in (1.1)—which we shall call *conic*—with inputs drawn from the uniform distribution on  $S^n$  that Demmel proved in [3] (shortly after [2]) a general result bounding their tail as a function of  $n$  and the degree of an algebraic hypersurface containing  $\Sigma$ .

Very recently, a new paradigm for probabilistic analysis was proposed by Spielman and Teng [6, 7]. Called *smoothed analysis*, it consists of replacing the idea of “random

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data” by that of “random perturbation of a given data” and study the worst-case (w.r.t. data  $a$ ) of the latter. In its original formulation, and in the case of a condition number  $\mathcal{C}(a)$ , this amounts to study the tail

$$\sup_{a \in \mathbb{R}^{n+1}} \text{Prob}_{z \in N(a, \sigma^2)} \{ \mathcal{C}(z) \geq t \}$$

or the expected value

$$\sup_{a \in \mathbb{R}^{n+1}} \mathbf{E}_{z \in N(a, \sigma^2)} [\ln \mathcal{C}(z)]$$

where  $N(a, \sigma^2)$  is a Gaussian distribution centered at  $a$  with covariance matrix  $\sigma^2 \text{Id}$  and  $\sigma^2$  small (with respect to  $\|a\|$ ). In [1], to obtain general results as in [3], data was again restricted to  $S^n$  and the expressions above replaced by

$$\sup_{a \in S^n} \text{Prob}_{z \in B(a, \sigma)} \{ \mathcal{C}(z) \geq t \}$$

and

$$\sup_{a \in S^n} \mathbf{E}_{z \in B(a, \sigma)} [\ln \mathcal{C}(z)]$$

where  $B(a, \sigma)$  is the open ball (that is, the spherical cap) in  $S^n$  centered at  $a$  and of radius  $\sigma$ , and  $z$  is drawn from a uniform distribution on this ball.

One of the claimed advantages of smoothed analysis is a smaller dependence on the underlying distribution. It follows from this claim that the replacement of Gaussian perturbations by uniform ones should not significantly affect the smoothed analysis of  $\mathcal{C}(a)$ . The goal of this note is to further pursue this claim by extending the main result in [1], combining it with ideas from [4], to a class of distributions we call *adversarial*. The support of such a distribution is, as in the uniform case, the ball  $B(a, \sigma)$  and they are radially symmetric as well. But their density increases when approaching  $a$  and has a pole at  $a$ .

**2. Preliminaries.** We assume our data space is  $\mathbb{R}^{n+1}$ , endowed with a scalar product  $\langle \cdot, \cdot \rangle$ . In all that follows we consider problems whose set of ill-posed inputs  $\Sigma$  is a point-symmetric cone in  $\mathbb{R}^{n+1}$ . That is, if  $x \in \Sigma$  then  $\lambda x \in \Sigma$  for all  $\lambda \in \mathbb{R}$ . By a *conic condition number* we understand a function  $\mathcal{C}: \mathbb{R}^{n+1} \rightarrow [1, \infty]$  such that for all  $a \in \mathbb{R}^{n+1}$  we have

$$\mathcal{C}(a) = \frac{\|a\|}{\text{dist}(a, \Sigma)},$$

where  $\| \cdot \|$  and  $\text{dist}$  are the norm and distance induced by  $\langle \cdot, \cdot \rangle$ . Note that for  $\lambda \neq 0$  we have  $\mathcal{C}(\lambda a) = \mathcal{C}(a)$ . We can therefore work with the  $n$ -dimensional real projective space  $\mathbb{P}^n$  as ambient space. If we also denote by  $\Sigma \subset \mathbb{P}^n$  the image of the ill-posed cone in projective space, then for  $a \in \mathbb{P}^n$  it follows that

$$\mathcal{C}(a) = \frac{1}{d_{\mathbb{P}}(a, \Sigma)},$$

where  $d_{\mathbb{P}}(x, y) = \sin \alpha$ , denotes the projective distance between  $x, y \in \mathbb{P}^n$  ( $\alpha$  being the angle between  $x$  and  $y$ ).

The two-fold covering  $p: S^n \rightarrow \mathbb{P}^n$  induces a measure  $\nu$  on  $\mathbb{P}^n$  by means of  $\nu(B) := \frac{1}{2} \text{Vol}_n(p^{-1}(B))$  for  $B \subseteq \mathbb{P}^n$ , where  $\text{Vol}_n$  is the  $n$ -dimensional volume on the sphere. Thus  $\nu(\mathbb{P}^n) = \mathcal{O}_n/2$ , where  $\mathcal{O}_n := \text{Vol}_n(S^n) = \frac{2\pi^{\frac{n+1}{2}}}{\Gamma(\frac{n+1}{2})}$ .

For  $0 < \sigma \leq 1$  we denote by  $B_{\mathbb{P}}(a, \sigma)$  the open ball of projective radius  $\sigma$  around  $a \in \mathbb{P}^n$ . It is known that

$$\nu(B_{\mathbb{P}}(a, \sigma)) = \mathcal{O}_{n-1} \cdot I_n(\sigma),$$

where

$$I_n(\sigma) := \int_0^\sigma \frac{r^{n-1}}{\sqrt{1-r^2}} dr. \quad (2.1)$$

The following bounds will prove useful on several occasions:

$$\frac{\sigma^n}{n} \leq I_n(\sigma) \leq \min \left\{ \frac{1}{\sqrt{1-\sigma^2}}, \sqrt{\frac{\pi n}{2}} \right\} \cdot \frac{\sigma^n}{n}. \quad (2.2)$$

For  $a \in \mathbb{P}^n$  and  $\sigma \in (0, 1]$  the uniform measure on  $B_{\mathbb{P}}(a, \sigma)$  is defined by

$$\nu_{a,\sigma}(B) = \frac{\nu(B \cap B_{\mathbb{P}}(a, \sigma))}{\nu(B_{\mathbb{P}}(a, \sigma))} \quad (2.3)$$

for all Borel-measurable  $B \subseteq \mathbb{P}^n$ .

**2.1. Uniform smoothed analysis.** A reformulation of the main result in [1] in the projective space setting can be written as follows.

**THEOREM 2.1.** *Let  $\mathcal{C}$  be a conic condition number with set of ill-posed inputs  $\Sigma \subset \mathbb{P}^n$ . Assume that  $\Sigma$  is contained in the zero set in  $\mathbb{P}^n$  of homogeneous polynomials of degree at most  $d$ . Then, for all  $\sigma \in (0, 1]$  and all  $t \geq t_0 = (2d+1)\frac{n}{\sigma}$ ,*

$$\sup_{a \in \mathbb{P}^n} \sup_{z \in B_{\mathbb{P}}(a, \sigma)} \text{Prob} \{ \mathcal{C}(z) \geq t \} \leq 13dn \frac{1}{\sigma t}.$$

and

$$\sup_{a \in \mathbb{P}^n} \sup_{z \in B_{\mathbb{P}}(a, \sigma)} \mathbf{E} [\ln \mathcal{C}(z)] \leq 2 \ln n + 2 \ln d + 2 \ln \frac{1}{\sigma} + 5,$$

where  $\text{Prob}$  and  $\mathbf{E}$  are taken with respect to  $\nu_{a,\sigma}$ .

As a consequence of this result, uniform smoothed analysis results for the condition numbers of a variety of problems are obtained, including linear equation solving, Moore-Penrose inversion, eigenvalue computation and polynomial system solving. The bounds obtained are consistently of the same order of magnitude as the best bounds obtained previously by ad-hoc methods.

**2.2. Uniformly Absolutely Continuous Distributions.** In [4] a general boosting mechanism was developed that allows extending any probabilistic analysis of a condition number with respect to some chosen probability distribution over the input data to a more general class of distributions.

Let  $\mu$  be a  $\nu_{a,\sigma}$ -absolutely continuous probability measure. Using the convention  $\ln(0) := -\infty$  we define, for  $\delta \in (0, 1)$ ,

$$\inf(\delta) := \inf \left\{ \frac{\ln \mu(B)}{\ln \nu_{a,\sigma}(B)} : B \text{ is Borel-measurable and } 0 < \nu_{a,\sigma}(B) \leq \delta \right\}$$

With these conventions, Theorem 2.2 of [4] shows that

$$\alpha_{\nu_{a,\sigma}}(\mu) := \liminf_{\delta \rightarrow 0}(\delta) \in [0, 1]. \quad (2.4)$$

Absolute continuity alone ensures that all  $\nu_{a,\sigma}$ -null-sets must be  $\mu$ -null-sets, but this does not imply that  $\mu(B)$  is small when  $\nu_{a,\sigma}(B)$  is small and strictly positive. In contrast, when  $\alpha_{\nu_{a,\sigma}}(\mu) > 0$  then (2.4) gives uniform upper bounds on  $\mu(B)$  in terms of  $\nu_{a,\sigma}(B)$ . Furthermore, the smaller  $\alpha$  gets, the larger the variation of  $\mu$  in terms of  $\nu_{a,\sigma}$ . If  $\mu$  is  $\nu_{a,\sigma}$ -absolutely continuous and  $\alpha_{\nu_{a,\sigma}}(\mu) > 0$ , we therefore say that  $\mu$  is *uniformly*  $\nu_{a,\sigma}$ -absolutely continuous and call  $\alpha_{\nu_{a,\sigma}}(\mu)$  the *smoothness parameter* of  $\mu$  with respect to  $\nu_{a,\sigma}$ .

The following result, which easily follows from (2.4), can be used to boost bounds on tail probabilities with respect to  $\nu_{a,\sigma}$  (as those in Theorem 2.1) to obtain similar bounds on any uniformly  $\nu_{a,\sigma}$ -absolutely continuous probability measure  $\mu$ .

**PROPOSITION 2.2.**  *$\alpha_{\nu_{a,\sigma}}(\mu)$  is the largest nonnegative real number  $\alpha$  for which it is true that for all  $\varepsilon > 0$  there exists  $\delta_\varepsilon > 0$  such that  $\nu_{a,\sigma}(B) \leq \delta_\varepsilon$  implies  $\mu(B) \leq \nu_{a,\sigma}(B)^{\alpha-\varepsilon}$ .*

**3. Smoothed analysis for adversarial distributions.** In this section we present our main result, namely an extension of Theorem 2.1 to the case where we have a radially symmetric distribution whose density has a pole at the point being perturbed. We begin by introducing some notation.

Let  $a \in \mathbb{P}^n$  and  $\sigma \in (0, 1]$ , and let  $\nu_{a,\sigma}$  be the uniform measure on  $B_{\mathbb{P}}(a, \sigma)$ , as defined in (2.3). Let  $\mu$  be a  $\nu_{a,\sigma}$ -absolutely continuous probability measure on  $\mathbb{P}^n$  with density  $f(x)$ . In other words,

$$\mu(B) = \int_B f(x) \nu_{a,\sigma}(dx)$$

for all events  $B$ . Assume further that  $f: \mathbb{P}^n \rightarrow [0, \infty]$  is of the form  $f(x) = g(d_{\mathbb{P}}(x, a))$ , with a monotonically decreasing function  $g: [0, \sigma] \rightarrow [0, \infty]$  of the form

$$g(r) = C_{\beta,\sigma} \cdot r^{-\beta} \cdot h(r),$$

with  $\beta < n$ , where  $C_{\beta,\sigma} = I_n(\sigma)/I_{n-\beta}(\sigma)$  and  $h: [0, \sigma] \rightarrow \mathbb{R}_+$  is a continuous function satisfying  $h(0) \neq 0$  and

$$\int_0^\sigma h(r) \frac{r^{n-\beta-1}}{\sqrt{1-r^2}} dr = I_{n-\beta}(\sigma),$$

so that  $\mu$  is a probability measure on  $B_{\mathbb{P}}(a, \sigma)$ . In other words,  $f$  is radially symmetric around  $a$  with respect to  $d_{\mathbb{P}}$  and has a pole of order  $-\beta$  at 0 in case  $\beta > 0$ . The normalizing factor  $C_{\beta,\sigma}$  is chosen to make  $h(r) = 1$  a valid choice. Set  $H := \sup_{0 \leq r \leq \sigma} h(r)$ . Note that  $H \geq 1$ , and that  $H = 1$  implies  $h \equiv 1$ .

It will be important to have expressions for  $\nu_{a,\sigma}(B)$  and  $\mu(B)$  when  $B = B_{\mathbb{P}}(a, \rho)$

is a projective ball. In this situation we have

$$\begin{aligned}
\mu(B_{\mathbb{P}}(a, \rho)) &= \frac{1}{\nu(B_{\mathbb{P}}(a, \sigma))} \int_{B_{\mathbb{P}}(a, \rho)} f(x) \nu(dx) \\
&= \frac{1}{\mathcal{O}_{n-1} I_n(\sigma)} \cdot C_{\beta, \sigma} \cdot \mathcal{O}_{n-1} \int_0^\rho r^{-\beta} h(r) \frac{r^{n-1}}{\sqrt{1-r^2}} dr \\
&= \frac{1}{I_{n-\beta}(\sigma)} \int_0^\rho h(r) \frac{r^{n-\beta-1}}{\sqrt{1-r^2}} dr \\
&\leq \left( \sup_{0 \leq r \leq \rho} h(r) \right) \cdot \frac{I_{n-\beta}(\rho)}{I_{n-\beta}(\sigma)}.
\end{aligned} \tag{3.1}$$

Similarly,

$$\mu(B_{\mathbb{P}}(a, \rho)) \geq \left( \inf_{0 \leq r \leq \rho} h(r) \right) \cdot \frac{I_{n-\beta}(\rho)}{I_{n-\beta}(\sigma)}.$$

In particular,

$$\nu_{a, \sigma}(B_{\mathbb{P}}(a, \rho)) = \frac{I_n(\rho)}{I_n(\sigma)}. \tag{3.2}$$

The main result of this note is the following.

**THEOREM 3.1.** *Let  $\mathcal{C}$  be a conic condition number with set of ill-posed inputs  $\Sigma \subseteq \mathbb{P}^n$ , and assume  $\Sigma$  is contained in a projective hypersurface of degree at most  $d$ . Then*

$$\mathbf{E}_{\mu}[\ln \mathcal{C}] \leq 2 \ln(n) + \ln(d) + \ln\left(\frac{1}{\sigma}\right) + \ln\left(\frac{13\pi}{2}\right) + \frac{1}{1 - \frac{\beta}{n}} \left( \ln \frac{2eH^2n}{\ln(\pi n/2)} \right).$$

This result applies to the variety of problems mentioned after Theorem 2.1. The statement of the Theorem follows from calculating the smoothness parameter  $\alpha_{\nu}(\mu)$  and the constants in Proposition 2.2. These are given by the following two lemmas, to be proven later.

**LEMMA 3.2.** *The smoothness parameter of  $\mu$  with respect to  $\nu_{a, \sigma}$  is given by  $\alpha_{\nu_{a, \sigma}}(\mu) = 1 - \beta/n$ .*

For the statement of the next Lemma, let  $\varepsilon \in (0, 1 - \beta/n)$ , and let

$$\rho_{\varepsilon} := \sigma \cdot \left( \frac{1}{H} \cdot \sqrt{1 - \left( \frac{2}{\pi n} \right)^{(1 - \frac{\beta}{n} - \varepsilon)/(n\varepsilon)}} \right)^{\frac{1}{\varepsilon n}} \left( \sqrt{\frac{2}{\pi n}} \right)^{(1 - \frac{\beta}{n} - \varepsilon) \frac{1}{\varepsilon n}}.$$

Set  $\delta_{\varepsilon} := I_n(\rho_{\varepsilon})/I_n(\sigma)$ .

**LEMMA 3.3.** *Let  $B \subseteq \mathbb{P}^n$  be such that  $\nu_{a, \sigma}(B) \leq \delta_{\varepsilon}$ . Then  $\mu(B) \leq (\nu_{a, \sigma}(B))^{1 - \frac{\beta}{n} - \varepsilon}$ .*

We are now ready to prove the main result.

PROOF OF THEOREM 3.1. Setting  $\varepsilon = \frac{1}{2}(1 - \frac{\beta}{n})$  and using the bounds (2.2) we obtain

$$\frac{2}{\pi n} \left( \frac{1}{H} \cdot \sqrt{1 - \left( \frac{2}{\pi n} \right)^{\frac{1}{n}}} \right)^{\frac{2}{1 - \frac{\beta}{n}}} \leq \delta_\varepsilon \leq \left( \frac{1}{H} \cdot \sqrt{1 - \left( \frac{2}{\pi n} \right)^{\frac{1}{n}}} \right)^{\frac{2}{1 - \frac{\beta}{n}}}. \quad (3.3)$$

From Theorem 2.1 it follows that for all  $t \geq t_0 := \ln[(1 + 2d)n/\sigma]$ ,

$$\text{Prob}_{\nu_{a,\sigma}}\{\ln \mathcal{C} > t\} \leq \frac{13dn}{\sigma} e^{-t}. \quad (3.4)$$

Set

$$t_\varepsilon := \ln \left( \frac{13dn}{\sigma \cdot \delta_\varepsilon} \right) = \ln \left( \frac{13dn}{\sigma} \right) + \ln(\delta_\varepsilon^{-1}).$$

Using (3.3) we obtain

$$\ln \left( 13 \frac{dn}{\sigma} \right) \leq t_\varepsilon - \frac{2}{1 - \frac{\beta}{n}} \ln \left( \frac{H}{\sqrt{1 - \left( \frac{2}{\pi n} \right)^{\frac{1}{n}}}} \right) \leq \ln \left( 13 \frac{\pi}{2} \frac{dn^2}{\sigma} \right).$$

The lower bound shows that  $t_\varepsilon > t_0$ , so that for all  $t \geq t_\varepsilon$ ,

$$\nu_{a,\sigma}(\{x : \ln \mathcal{C}(x) > t\}) = \text{Prob}_{\nu_{a,\sigma}}\{\ln \mathcal{C} > t\} \leq \frac{13dn}{\sigma} e^{-t} \leq \delta_\varepsilon.$$

Applying Lemma 3.3, it follows that for  $t \geq t_\varepsilon$ ,

$$\text{Prob}_\mu\{\ln \mathcal{C} > t\} = \mu(\{x : \ln \mathcal{C}(x) > t\}) \leq \left( \frac{13dn}{\sigma} e^{-t} \right)^{\frac{1}{2}(1 - \frac{\beta}{n})},$$

and hence,

$$\begin{aligned} \mathbf{E}_\mu[\ln \mathcal{C}] &= \int_0^\infty \text{Prob}_\mu\{\ln \mathcal{C} > t\} dt \\ &\leq \int_0^{t_\varepsilon} 1 dt + \int_{t_\varepsilon}^\infty \left( \frac{13dn}{\sigma} e^{-t} \right)^{\frac{1}{2}(1 - \frac{\beta}{n})} dt \\ &= t_\varepsilon + \frac{2\delta_\varepsilon^{\frac{1}{2}(1 - \frac{\beta}{n})}}{1 - \frac{\beta}{n}}. \end{aligned}$$

Using the bounds on  $t_\varepsilon$  and  $\delta_\varepsilon$  we get

$$\mathbf{E}_\mu[\ln \mathcal{C}] \leq 2 \ln(n) + \ln(d) + \ln \left( \frac{1}{\sigma} \right) + \ln \left( \frac{13\pi}{2} \right) + \frac{2}{1 - \frac{\beta}{n}} \left( \ln \left( \frac{H}{\sqrt{1 - \left( \frac{2}{\pi n} \right)^{\frac{1}{n}}}} \right) + \frac{\sqrt{1 - \left( \frac{2}{\pi n} \right)^{\frac{1}{n}}}}{H} \right).$$

A small calculation shows that  $\left( 1 - \left( \frac{2}{\pi n} \right)^{\frac{1}{n}} \right)^{-1/2} \leq \sqrt{\frac{2n}{\ln(\pi n/2)}}$ . This completes the proof.  $\square$

**3.1. Proofs of Lemmas 3.2 and 3.3.** The content of the following Lemma, needed for calculating the smoothness parameter, should be intuitively clear.

LEMMA 3.4. *Let  $0 < \delta < 1$ . Then among all measurable sets  $B \subseteq B_{\mathbb{P}}(a, \sigma)$  with  $0 < \nu_{a, \sigma}(B) \leq \delta$ ,  $\mu(B)$  is maximized by  $B_{\mathbb{P}}(a, \rho)$  where  $\rho \in (0, \sigma)$  is chosen so that  $\nu_{a, \sigma}(B_{\mathbb{P}}(a, \rho)) = \delta$ .*

PROOF. It clearly suffices to show that

$$\int_B f(x) \nu_{a, \sigma}(dx) \leq \int_{B_{\mathbb{P}}(a, \rho)} f(x) \nu_{a, \sigma}(dx)$$

for all Borel sets  $B \subset B_{\mathbb{P}}(a, \sigma)$  such that  $\nu_{a, \sigma}(B) = \delta$ . Indeed, we have

$$\begin{aligned} \int_B f(x) \nu_{a, \sigma}(dx) &= \int_{B \cap B_{\mathbb{P}}(a, \rho)} f(x) \nu_{a, \sigma}(dx) + \int_{B \setminus B_{\mathbb{P}}(a, \rho)} f(x) \nu_{a, \sigma}(dx) \\ &\leq \int_{B \cap B_{\mathbb{P}}(a, \rho)} f(x) \nu_{a, \sigma}(dx) + g(\rho) \nu_{a, \sigma}(B \setminus B_{\mathbb{P}}(a, \rho)) \\ &= \int_{B \cap B_{\mathbb{P}}(a, \rho)} f(x) \nu_{a, \sigma}(dx) + g(\rho) \nu_{a, \sigma}(B_{\mathbb{P}}(a, \rho) \setminus B) \\ &\leq \int_{B \cap B_{\mathbb{P}}(a, \rho)} f(x) \nu_{a, \sigma}(dx) + \int_{B_{\mathbb{P}}(a, \rho) \setminus B} f(x) \nu_{a, \sigma}(dx) \\ &= \int_{B_{\mathbb{P}}(a, \rho)} f(x) \nu_{a, \sigma}(dx), \end{aligned} \tag{3.5}$$

where we have used  $\nu_{a, \sigma}(B_{\mathbb{P}}(a, \rho)) = \delta = \nu_{a, \sigma}(B)$  in (3.5). This proves our claim.  $\square$

Even though  $\rho$  is a function of  $\delta$ , we will not reflect this notationally in the sequel.

PROOF OF LEMMA 3.2. From (3.1), (3.2) and (2.2) we get the bounds of the form

$$\frac{1}{C_1} \cdot \rho^n \leq \nu_{a, \sigma}(B_{\mathbb{P}}(a, \rho)) \leq C_1 \cdot \rho^n, \tag{3.6}$$

$$\inf_{0 \leq r \leq \rho} h(r) \cdot \frac{1}{C_2} \cdot \rho^{n-\beta} \leq \mu(B_{\mathbb{P}}(a, \rho)) \leq \sup_{0 \leq r \leq \rho} h(r) \cdot C_2 \cdot \rho^{n-\beta}, \tag{3.7}$$

where the constants  $C_i$  do not depend on  $\rho$ .

We thus have (using Lemma 3.4)

$$\begin{aligned} \alpha_{\nu_{a, \sigma}}(\mu) &= \liminf_{\delta \rightarrow 0} \left\{ \frac{\ln \mu(B)}{\ln \nu_{a, \sigma}(B)} : B \text{ measurable, } 0 < \nu_{a, \sigma}(B) \leq \delta \right\} \\ &= \lim_{\rho \rightarrow 0} \frac{\ln \mu(B_{\mathbb{P}}(a, \rho))}{\ln \nu_{a, \sigma}(B_{\mathbb{P}}(a, \rho))} \\ &\begin{cases} \leq \lim_{\rho \rightarrow 0} \frac{\ln(\inf h(r)/C_2) + (n-\beta) \ln \rho}{\ln(C_1) + n \ln \rho} = 1 - \frac{\beta}{n} \\ \geq \lim_{\rho \rightarrow 0} \frac{\ln(C_2 \cdot \sup h(r)) + (n-\beta) \ln \rho}{-\ln C_1 + n \ln \rho} = 1 - \frac{\beta}{n}. \end{cases} \end{aligned}$$

This concludes the proof.  $\square$



PROOF OF LEMMA 3.3. Since sets of the form  $B_{\mathbb{P}}(a, \rho)$  maximise  $\mu(B)$  among all measurable sets  $B \subseteq B_{\mathbb{P}}(a, \sigma)$  such that  $\nu_{a, \sigma}(B) \leq \delta$  for any  $\delta$ , we may w.l.o.g. assume  $B = B_{\mathbb{P}}(a, \rho)$ . By (3.1) and (3.2) our task amounts to showing

$$H \cdot \frac{I_{n-\beta}(\rho)}{I_{n-\beta}(\sigma)} \leq \left( \frac{I_n(\rho)}{I_n(\sigma)} \right)^{1-\frac{\beta}{n}-\varepsilon}$$

for  $\rho \leq \rho_\varepsilon$ . And indeed, using the bounds (2.2), we get

$$\begin{aligned} H \cdot \frac{I_{n-\beta}(\rho)}{I_{n-\beta}(\sigma)} &\leq H \frac{1}{\sqrt{1-\rho^2}} \cdot \left( \frac{\rho}{\sigma} \right)^{n-\beta} \\ &\leq H \frac{1}{\sqrt{1-\rho^2}} \cdot \left( \left( \frac{\rho}{\sigma} \right)^n \right)^{1-\frac{\beta}{n}-\varepsilon} \left( \frac{\rho_\varepsilon}{\sigma} \right)^{\varepsilon n} \\ &\leq \frac{\sqrt{1 - \left( \frac{2}{\pi n} \right)^{(1-\frac{\beta}{n}-\varepsilon)/(n\varepsilon)}}}{\sqrt{1-\rho^2}} \cdot \left( \sqrt{\frac{2}{\pi n}} \left( \frac{\rho}{\sigma} \right)^n \right)^{1-\frac{\beta}{n}-\varepsilon} \\ &\leq \frac{\sqrt{1 - \left( \frac{2}{\pi n} \right)^{(1-\frac{\beta}{n}-\varepsilon)/(n\varepsilon)}}}{\sqrt{1-\rho^2}} \cdot \left( \frac{I_n(\rho)}{I_n(\sigma)} \right)^{1-\frac{\beta}{n}-\varepsilon}, \end{aligned}$$

where for the last inequality we use the bounds (2.2) again. Moreover, we have

$$\rho \leq \rho_\varepsilon \leq \left( \sqrt{\frac{2}{\pi n}} \right)^{(1-\frac{\beta}{n}-\varepsilon)\frac{1}{\varepsilon n}}.$$

Therefore,  $\sqrt{1 - \left( \frac{2}{\pi n} \right)^{(1-\frac{\beta}{n}-\varepsilon)\frac{1}{\varepsilon n}}} \leq \sqrt{1-\rho^2}$ , completing the proof.  $\square$

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